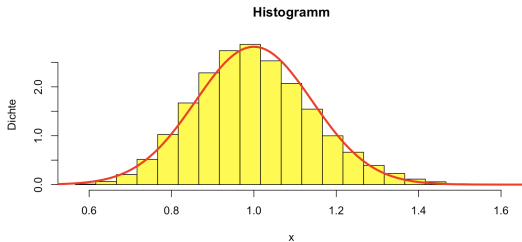


Central limit theorem and Law of large numbers



Reminder

- $X \sim \mathcal{N}(\mu, \sigma^2)$
 - One of the most important distributions in statistics and probability theory.



Normal distribution on the front of the German 10 Mark banknote from 1990s

- X has a **Normal distribution** (or **Gauss distribution**) with the parameters $\mu \in \mathbb{R}$ and $\sigma > 0$, if its pdf is given by

$$f(x) = \frac{1}{\sqrt{2\pi} \sigma} \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}} \quad \text{for } x \in \mathbb{R}$$

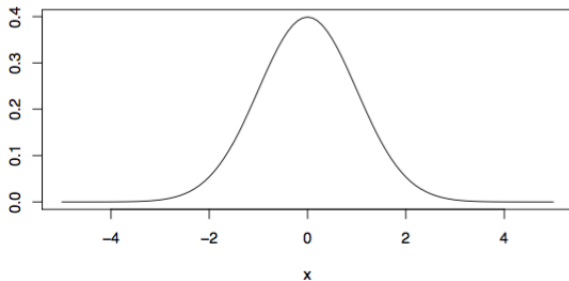
- $\mathbb{E}(X) = \mu, \quad \text{Var}(X) = \sigma^2$

$$\underline{Y = aX + b \sim \mathcal{N}(a\mu + b, a^2\sigma^2)}$$

We recall

- $Z \sim \mathcal{N}(0,1)$ is the **Standard normal random variable**

... $\mathbb{E}(Z) = 0$, $\text{Var}(Z) = 1$



- Its cdf is given by $\Phi(z) = P(Z \leq z)$
- The following hold

$$P(-1 \leq Z \leq 1) \approx 0.68$$

$$P(-2 \leq Z \leq 2) \approx 0.95$$

$$P(-3 \leq Z \leq 3) \approx 0.997$$

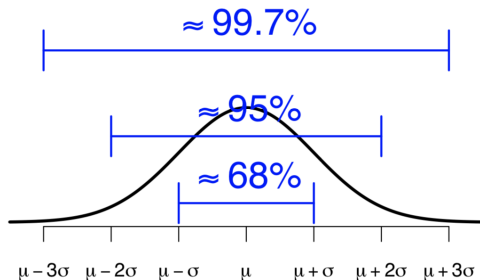
We recall

- Let $X \sim \mathcal{N}(\mu, \sigma^2)$.

- Standardization

$$Z = \frac{X - \mu}{\sigma} \sim \mathcal{N}(0, 1)$$

- For $X \sim \mathcal{N}(\mu, \sigma^2)$ the 68-95-99.7-Rule holds



Sample mean

- Let X_1, X_2, \dots, X_n be **independent identically distributed (i.i.d.)** random variables with $\mathbb{E}(X_1) = \mu$ and $\text{Var}(X_1) = \sigma^2$.

... X_1, X_2, \dots, X_n is also called a **random sample of size n**

- For the **sample mean** $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ it holds

$$\mathbb{E}(\bar{X}) = \mu \quad \text{and} \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n}$$

$$\begin{aligned} \bullet \mathbb{E}(\bar{X}) &= \mathbb{E}\left(\frac{1}{n} \sum_{i=1}^n X_i\right) \stackrel{\text{linearity}}{=} \frac{1}{n} \sum_{i=1}^n \mathbb{E}(X_i) \stackrel{\text{ident. distr.}}{=} \frac{1}{n} \sum_{i=1}^n \mathbb{E}(X_1) \\ &= \frac{1}{n} \cdot n \mathbb{E}(X_1) = \mathbb{E}(X_1) = \mu \end{aligned}$$

$$\begin{aligned} \bullet \text{Var}(\bar{X}) &= \text{Var}\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n^2} \text{Var}\left(\sum_{i=1}^n X_i\right) \stackrel{\text{independ.}}{=} \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) \\ &\stackrel{\text{ident. distr.}}{=} \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_1) = \frac{1}{n^2} \cdot n \text{Var}(X_1) = \frac{\text{Var}(X_1)}{n} = \frac{\sigma^2}{n} \end{aligned}$$

Sample mean

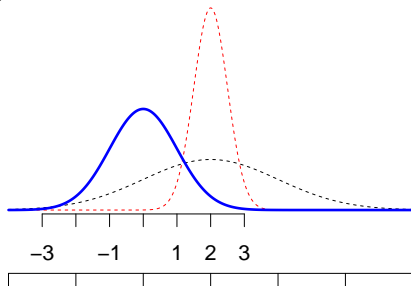
- Let X_1, X_2, \dots, X_n be independent identically distributed (i.i.d.) random variables with $E(X_1) = \mu$ and $Var(X_1) = \sigma^2$.

... X_1, X_2, \dots, X_n is also called a random sample of size n

- For the sample mean $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ it holds

$$E(\bar{X}) = \mu \quad \text{and} \quad Var(\bar{X}) = \frac{\sigma^2}{n}$$

- Especially if X_1, \dots, X_n are normally distributed, the sample mean is also normally distributed $\bar{X} \sim \mathcal{N}(\mu, \frac{\sigma^2}{n})$



Convergence

- Two important theorems of probability theory:
 - Law of Large Numbers (LLN)
 - in German: Gesetz der großen Zahlen (GGZ)
 - Central Limit Theorem (CLT)
 - in German: Zentraler Grenzwertsatz (ZGWS)

... We state the theorems for
independent and identically distributed random variables

Theorems

- Let X_1, \dots, X_n be independent and identically distributed (i.i.d.) random variables with $\mathbb{E}(X_1) = \mu$ and finite $\text{Var}(X_1) = \sigma^2 < \infty$.

... n i.i.d. random variables X_1, \dots, X_n are called a random sample of size n

- Note that X_1, \dots, X_n are not necessarily normally distributed.
- We consider the sample mean

$$\bar{X}_n = \frac{X_1 + \dots + X_n}{n} = \frac{1}{n} \sum_{i=1}^n X_i$$

- The Law of large numbers and the Central limit theorem provide information about the value and the distribution of \bar{X}_n .
 - LLN: As n grows, the probability of \bar{X}_n being in the neighborhood of μ tends to 1.
 - CLT: For large n , the distribution of \bar{X}_n is approximately $\mathcal{N}(\mu, \frac{\sigma^2}{n})$.

Law of large numbers

- The Law of large numbers says that the sample mean \bar{X}_n will be with high probability very close to the expectation μ of the underlying distribution
- Formal statement:

Suppose $X_1, X_2, \dots, X_n \dots$ are i.i.d. random variables with expectation μ and finite variance σ^2 . For each n , let \bar{X}_n be the mean of the first n variables. Then for any $a > 0$, we have

$$\lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| < a) = 1$$

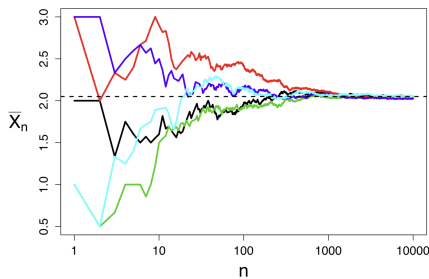
Example

- We consider a discrete random variable with the pmf

x	0	1	2	3	4
$p(x)$	0.1	0.2	0.3	0.35	0.05

- Then, $\mathbb{E}(X) = \sum_{x=0}^4 x p(x) = 2.05$ and $\text{Var}(X) = \sum_{x=0}^4 (x - 2.05)^2 p(x) = 1.1475$
- Since the variance of X exists, the Law of Large Numbers says

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \rightarrow \mathbb{E}(X) = 2.05 \quad \text{for large } n$$



Example 2

- The mean values of the Bernoulli random variables
 - Let X_1, \dots, X_n denote n independent flips of a fair coin
 - $X_i \sim \text{ber}(0.5)$ and $\mathbb{E}(X_i) = \mu = 0.5$
 - $S_n = X_1 + \dots + X_n \sim B(n, 0.5)$
 - Then $\bar{X}_n = \frac{1}{n}S_n$ is the **average number** of heads obtained in n flips.
 - We expect that this **number** for large n is close to $\mu = 0.5$.
 - **LLN**: The sample mean \bar{X}_n is very likely to be close to the expected value **0.5** for large n .
 - For example, let us consider the probability of being within **0.1** from the expected value $\mu = 0.5$, i.e.
$$P(|\bar{X}_n - \mu| \leq 0.1) \text{ or equivalently } P(0.4 \leq \bar{X}_n \leq 0.6) = P(0.4n \leq S_n \leq 0.6n)$$
 - in R:

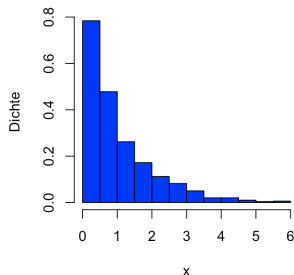
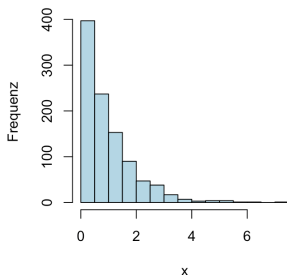
$n = 10$:	<code>diff(pbinom(c(3,6), 10, 0.5))</code>	<code>= 0.65625</code>
$n = 50$:	<code>diff(pbinom(c(19,30), 50, 0.5))</code>	<code>= 0.8810795</code>
$n = 100$:	<code>diff(pbinom(c(39,60), 100, 0.5))</code>	<code>= 0.9647998</code>
$n = 500$:	<code>diff(pbinom(c(199,300), 500, 0.5))</code>	<code>= 0.9999941</code>
$n = 1000$:	<code>diff(pbinom(c(399,600), 1000, 0.5))</code>	<code>= 1</code>

HW Repeat the calculations and determine the probability

$$P(|\bar{X}_n - \mu| \leq 0.01).$$

Histograms

- Histograms (**frequency** and **density** histogram)
 - ... A graphical representation of a frequency distribution based on a previous classification of the data.
- Made by *binning* data ... data binning is a way to group a number of continuous values into a smaller number of *bins*
- **Frequency:** **height** of bar over bin = number of data points in bin.
- **Density:** **area** of bar is the fraction of all data points that lie in the bin.
 - ... Thus, total area is 1.



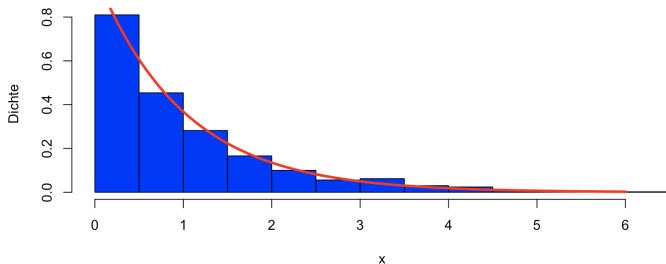
in R:

`hist(x, freq=TRUE)`

`hist(x, freq=FALSE)` or `hist(x, prob=TRUE)`

Law of Large Numbers and histograms

- The Law of Large Numbers implies that density histogram converges to probability density function.



The histogram with bin width 0.5 showing 1000 draws from an exponential $\exp(1)$ distribution. The pdf of $\exp(1)$ is given in red.

Central Limit Theorem

- The **Central Limit Theorem** states that the **sum**, resp. **mean**, of many independent copies of a random variable is approximately a **normal random variable**.
- Let X_1, X_2, \dots be a sequence of i.i.d. with expectation μ and standard deviation σ .
 - For each n :
 - the sum $S_n = X_1 + X_2 + \dots + X_n$
 - the sample mean $\bar{X}_n = \frac{1}{n}(X_1 + X_2 + \dots + X_n)$
 - The **Central Limit Theorem** says that for large n it holds

$$S_n \approx \mathcal{N}(n\mu, n\sigma^2)$$

$$\bar{X}_n \approx \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$$

- Thus, the standardized \bar{X}_n and S_n have approximately $\mathcal{N}(0,1)$, i.e.

$$\frac{S_n - n\mu}{\sqrt{n}\sigma} \approx \mathcal{N}(0,1) \quad \text{and} \quad \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} \approx \mathcal{N}(0,1)$$

Example

- Simulation for CLT

- Let X_1, X_2, \dots be a sequence of i.i.d. $\exp(1)$ random variables.

- Then, $\mathbb{E}(X_i) = 1$ and $\text{Var}(X_i) = 1 < \infty$.

- We consider sample means $\bar{X}_n = \frac{1}{n}(X_1 + \dots + X_n)$.

- Central Limit Theorem:

- \bar{X}_n : $\mathbb{E}(\bar{X}_n) = 1$ and $\text{Var}(\bar{X}_n) = \frac{1}{n}$

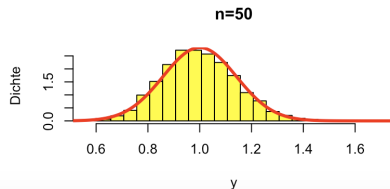
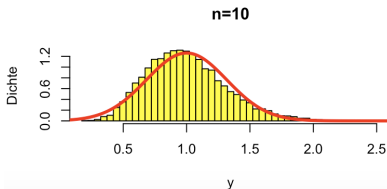
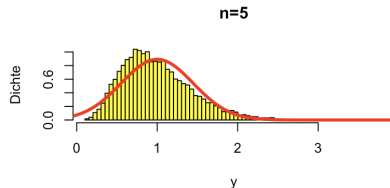
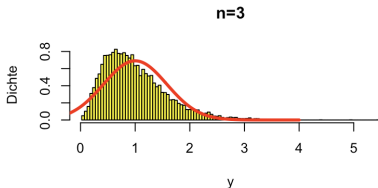
$$\bar{X}_n \approx \mathcal{N}(1, \frac{1}{n})$$

- The standardized means Y_n

$$Y_n = \frac{\bar{X}_n - \mathbb{E}(\bar{X}_n)}{\sqrt{\text{Var}(\bar{X}_n)}} = \frac{\bar{X}_n - 1}{\frac{1}{\sqrt{n}}} \approx \mathcal{N}(0, 1)$$

Example

- The following plots show the results for respectively $n = 3, 5, 10, 50$ based on 10000 simulated values for \bar{X}_n . The density of the $\mathcal{N}(1, \frac{1}{n})$ distribution is drawn in red.



Normal approximation of $B(n, p)$ distribution

- The **sum** $S_n = X_1 + \dots + X_n$ of independent $ber(p)$ random variables X_1, \dots, X_n has $B(n, p)$ distribution. $\mathbb{E}(S_n) = np$ and $\mathbb{Var}(S_n) = np(1-p)$
- Using the **continuity correction** it holds for $a \leq b$, where $a, b \in \{0, 1, \dots, n\}$:

$$P(a \leq S_n \leq b) \approx \Phi\left(\frac{b + \frac{1}{2} - np}{\sqrt{np(1-p)}}\right) - \Phi\left(\frac{a - \frac{1}{2} - np}{\sqrt{np(1-p)}}\right)$$

- Let T be a continuous random variable such that $\mathbb{E}(T) = \mathbb{E}(S_n) = np$ and $\mathbb{Var}(T) = \mathbb{Var}(S_n) = np(1-p)$. Then,

$$\begin{aligned} P(a \leq S_n \leq b) &= P\left(a - \frac{1}{2} \leq S_n < b + \frac{1}{2}\right) \approx P\left(a - \frac{1}{2} \leq T < b + \frac{1}{2}\right) \\ &= P\left(\frac{a - \frac{1}{2} - \mathbb{E}(T)}{\sqrt{\mathbb{Var}(T)}} \leq \frac{T - \mathbb{E}(T)}{\sqrt{\mathbb{Var}(T)}} < \frac{b + \frac{1}{2} - \mathbb{E}(T)}{\sqrt{\mathbb{Var}(T)}}\right) \\ &\stackrel{\text{CLT}}{\approx} P\left(\frac{a - \frac{1}{2} - np}{\sqrt{np(1-p)}} \leq Z < \frac{b + \frac{1}{2} - np}{\sqrt{np(1-p)}}\right) \\ &= \Phi\left(\frac{b + \frac{1}{2} - np}{\sqrt{np(1-p)}}\right) - \Phi\left(\frac{a - \frac{1}{2} - np}{\sqrt{np(1-p)}}\right) \end{aligned}$$

Normal approximation of $B(n, p)$ distribution

- The sum $S_n = X_1 + \dots + X_n$ of independent $ber(p)$ random variables X_1, \dots, X_n has $B(n, p)$ distribution.
 - $\mathbb{E}(S_n) = np$ and $\mathbb{V}ar(S_n) = np(1 - p)$
- Using the **continuity correction** it holds for $a \leq b$, where $a, b \in \{0, 1, \dots, n\}$:

$$P(a \leq S_n \leq b) \approx \Phi\left(\frac{b + \frac{1}{2} - np}{\sqrt{np(1-p)}}\right) - \Phi\left(\frac{a - \frac{1}{2} - np}{\sqrt{np(1-p)}}\right)$$

- Approximation of the cdf of S_n :

$$P(S_n \leq x) \approx \Phi\left(\frac{x + \frac{1}{2} - np}{\sqrt{np(1-p)}}\right)$$

Φ is the cdf of $\mathcal{N}(0, 1)$

- Rule of thumb: The approximation is considered reasonable, when $\min\{np, np(1-p)\} \geq 10$

Normal approximation of $\mathcal{P}(\lambda)$ distribution

- Let $X \sim \mathcal{P}(\lambda)$. Then, by providing the **continuity correction** for $a \leq b$, with $a, b \in \mathbb{N}_0$

$$P(a \leq X \leq b) \approx \Phi\left(\frac{b + \frac{1}{2} - \lambda}{\sqrt{\lambda}}\right) - \Phi\left(\frac{a - \frac{1}{2} - \lambda}{\sqrt{\lambda}}\right),$$

- Approximation of the cdf of X :

$$P(X \leq x) \approx \Phi\left(\frac{x + \frac{1}{2} - \lambda}{\sqrt{\lambda}}\right)$$

- Rule of thumb: The approximation is considered to be reasonable when $\lambda > 15$.

Example

- An accountant rounds to the nearest euro. We assume the error in rounding follows uniform distribution on $(-0.5, 0.5)$. Estimate the probability that the total error in 300 entries is more than 5 euro.

- Let X_j be the rounding error of the j th entry. Then, $X_j \sim U(-0.5, 0.5)$

$$\mathbb{E}(X_j) = 0 \text{ and } \text{Var}(X_j) = \frac{1}{12}$$

- The overall error is $S = X_1 + \dots + X_{300}$.

$$\mathbb{E}(S) = 0, \text{Var}(S) = \frac{300}{12} = 25 \text{ and } \sigma_S = 5.$$

- CLT: for the standardized S it holds

$$Z = \frac{S - \mathbb{E}(S)}{\sqrt{\text{Var}(S)}} = \frac{S}{5} \sim \mathcal{N}(0, 1)$$

- Then,

$$P(|S| > 5) = 1 - P(|S| \leq 5) \approx 1 - P(|Z| \leq 1) \approx 0.32$$

Computations:

The pdf is of the form $f(x) = 1$ if $x \in (-0.5, 0.5)$ and $f(x) = 0$ otherwise. Then, $\mathbb{E}(X_j) = \int_{-0.5}^{0.5} x dx = 0$ and

$$\text{Var}(X_j) = \mathbb{E}(X_j^2) = \int_{-0.5}^{0.5} x^2 dx = 2 \int_0^{0.5} x^2 dx = \frac{2}{3} x^3 \Big|_0^{0.5} = \frac{1}{12}. \text{ Thus, } \mathbb{E}S = 300 \cdot \mathbb{E}X_1 = 0 \text{ and } \text{Var}S = 300 \cdot \text{Var}(X_1) = 300 \cdot \frac{1}{12} = 25.$$

Examples

HW Let X_1, X_2, \dots, X_{25} be independent and identically distributed (i.i.d.) random variables and $X_1 \sim \mathcal{N}(1, 4)$. Find the probability $P(X_1 + X_2 + \dots + X_{25} \geq 26)$.

HW Transportation officials tell us that 60% of the population wear their seatbelts while driving. A random sample of 1000 drivers has been taken. What is the probability that between 580 and 630 of the drivers were wearing their seatbelts?

HW Let X_1, X_2, \dots, X_{121} be i.i.d. with the expectation $\mu = 35$ and variance $\sigma^2 = 25$. Denote by

$$\bar{X}_{121} = \frac{1}{121}(X_1 + \dots + X_{121})$$

the sample mean. Approximate the probability $P(\bar{X}_{121} > 35.2)$ using the Central limit theorem.

HW We toss a fair coin 100 times.
What is the probability of obtaining 60 or more heads?

A few multiple-choice questions

- (1) Let X_1, X_2, \dots, X_{81} be i.i.d. sample from a population with population mean $\mu = 5$ and population variance $\sigma^2 = 4$ and let $S = X_1 + X_2 + \dots + X_{81}$. Approximate the probability $P(S \notin [387, 423])$ using the Central limit theorem.
- a. 68%
 - b. 78%
 - c. 45%
 - d. 32%
- (2) Assume that X is a binomial random variable with $n = 100$ and $p = 0.1$. Use the normal probability distribution to compute $P(X \leq 15)$.
- a. 0.5336
 - b. 0.9664
 - c. 0.0336
 - d. 0.4664

A few multiple-choice questions

- (3) Transportation officials tell us that 60% of the population wear their seatbelts while driving. A random sample of 1000 drivers has been taken. What is the probability that between 220 and 550 of the drivers were wearing their seatbelts?
- a. 1.0
 - b. 0.4015
 - c. 0.9066
 - d. 0
- (4) Consider a sequence of independent tosses of a coin that is biased so that it comes up heads with probability 0.7 and tails with probability 0.3. Let X_i be 1 if the i th toss comes up heads and 0 otherwise. Let Y be the number of heads in the first 4900 tosses of the biased coin, i.e. $Y = \sum_{i=1}^{4900} X_i$. Use a normal random variable to approximate the probability that $Y \leq 3450$. Approximate the result in terms of the cdf $\Phi(z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-\frac{x^2}{2}} dx$.
- a. $\Phi(0.62)$
 - b. $\Phi(0.02)$
 - c. $1 - \Phi(1.41)$
 - d. $\Phi(1.92)$

A few multiple-choice questions

- (5) Let X_1, \dots, X_{64} be a random sample from a distribution with the expectation -1.2 and variance 4. Let

$$\bar{X} = \frac{1}{64} \sum_{i=1}^{64} X_i$$

be the sample mean. Determine the approximate value of $P(\bar{X} > -0.9)$ using the Central limit theorem and express it in terms of a suitable R-function.

- a. `pnorm(1.2)`
 - b. `pnorm(-1.2)`
 - c. `pnorm(-0.9, 1.2, 0.5)`
 - d. `pnorm(-0.9, -1.2, 0.25)`
- (6) The central limit theorem is important in statistics because
- a. For large n , it says that the population is approximately normal
 - b. For any population, it says the sampling distribution of the sample mean is approximately normal, regardless of the sample size.
 - c. For large n , it says the sampling distribution of the sample mean is approximately normal, regardless of the population.
 - d. For any size sample, it says that the sampling distribution of the sample mean is approximately normal.

A few multiple-choice questions

- (7) Consider a sequence of independent tosses of a coin that is biased so that it comes up heads with probability $3/4$ and tails with probability $1/4$. Let X be the number of heads in the first 4800 tosses of the biased coin. Use a normal random variable to approximate the probability

$$P(X > 3590).$$

In the computations include the continuity correction.

- a. 0.3632
 - b. 0.6293
 - c. 0.6368
 - d. 0.3707
- (8) If the heights of women are normally distributed with expectation 163 centimeters, which of the following is the smallest?
- a. The probability of randomly choosing one woman and finding her height is between 160 and 166 centimeters.
 - b. The probability of randomly choosing 15 women and finding that their mean height is between 160 and 166 centimeters.
 - c. The probability of randomly choosing 100 women and finding that their mean height is between 160 and 166 centimeters.
 - d. All of the above have the same probability.

A few multiple-choice questions

- (9) Assume that a baseball team has an average pitcher, that is, one whose probability of winning any decision is 0.5. If this pitcher has 30 decisions in a season, what is the probability that he will win at least 23 games?
- a. less than 0.5%
 - b. between 0.5% and 1%
 - c. between 1% and 5%
 - d. more than 5%
- (10) Assume the given distribution is normal. Pumpkins grown on a certain farm have weights with a mean of 10 kilograms. What is the standard deviation (in kilograms) of the weight if 37% of the pumpkins weigh more than 13 kilograms?
- a. 1.22
 - b. 2.97
 - c. 4.76
 - d. 8.82

Thank you for your attention!