

$\sum_i n_i = n$ absolute Häufigkeit n_i $f_i = \frac{n_i}{n}$ relativen Häufigkeiten	$x_{(1)} \leq x_{(2)} \leq \dots \leq x_{(n-1)} \leq x_{(n)}$ range = $x_{(n)} - x_{(1)}$ (Spannweite)	Lower bound = $Q_1 - 1.5 \cdot IQR$ Upper bound = $Q_3 + 1.5 \cdot IQR$ Interquartile range IQR = $Q_3 - Q_1$	
Median = $\begin{cases} x_{(k+1)}, & n = 2k + 1 \\ \frac{1}{2}(x_{(k)} + x_{(k+1)}), & n = 2k, \end{cases}$ Sample mean (Mittelwert) $\bar{x}_n = \frac{x_1 + x_2 + \dots + x_n}{n} = \frac{1}{n} \sum_{i=1}^n x_i$ Geometric mean (geometrische Mittel) $\bar{x}_n^{(g)} = \sqrt[n]{x_1 \cdot x_2 \cdot \dots \cdot x_n}$ Harmonic mean (harmonische Mittel) $\bar{x}_n^{(h)} = \frac{n}{\frac{1}{x_1} + \frac{1}{x_2} + \dots + \frac{1}{x_n}}$	Empirical central moment of order r $m_r = \frac{1}{n} \cdot \sum_{i=1}^n (x_i - \bar{x}_n)^r$ Sample variance (Stichprobenstreuung) $s_n^2 = \frac{1}{n-1} \cdot \sum_{i=1}^n (x_i - \bar{x}_n)^2 = \frac{1}{n-1} \left(\sum_{i=1}^n x_i^2 - n \bar{x}_n^2 \right)$ Correlation coefficient $-1 \leq r_{xy} \leq 1$ $r_{xy} = \frac{s_{xy}}{s_x \cdot s_y} = \frac{1}{n-1} \cdot \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s_x} \right) \cdot \left(\frac{y_i - \bar{y}}{s_y} \right)$ empirical covariance is $s_{xy} = \frac{1}{n-1} \cdot \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})$	Skewness $g_1^{(1)} = \frac{m_3}{m_2^{3/2}} = \frac{\sqrt{n}}{(n-1)^{3/2}} \cdot \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s} \right)^3$ $g_1^{(2)} = \frac{n}{(n-1)(n-2)} \cdot \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s} \right)^3$ $g_1^{(3)} = \frac{m_3}{s^3} = \frac{1}{n} \cdot \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s} \right)^3$ $g_1 > 0$ negative skewed : $mode < median < \bar{x}$ $g_1 < 0$ positive skewed : $mode > median > \bar{x}$	
Laplace space For $\Omega = \{\omega_1, \omega_2, \dots, \omega_n\}$, $ \Omega = n$ $P(A) = \frac{ A }{ \Omega } = \frac{\text{number of elements of } A}{\text{number of elements of } \Omega} \quad p_i = P(\{\omega_i\}) = \frac{1}{ \Omega } = \frac{1}{n}$	Sample space (Merkmalraum) Ω Event (Ereignis) $A \subseteq \Omega \quad P(A) = \sum_{\omega \in A} P(\omega)$ Measurable space (Messraum) (Ω, \mathcal{A}) Probability space (Wahrscheinlichkeitsraum) (Ω, \mathcal{A}, P)	Kurtosis $g_2^{(1)} = \frac{m_4}{m_2^2} = \frac{\sqrt{n}}{(n-1)^2} \cdot \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s} \right)^4$ $g_2^{(2)} = \frac{n(n+1)}{(n-1)(n-2)(n-3)} \cdot \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s} \right)^4$ $g_2^{(3)} = \frac{m_4}{s^4} = \frac{1}{n} \cdot \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s} \right)^4$	
Geometric probabilities $P(A) = \frac{\text{area of desired outcomes}}{\text{area of total outcomes}}$	Conditional probability $P(A B) = \frac{P(A \cap B)}{P(B)}$	rules of probability (1) $P(A^c) = 1 - P(A)$ (2) If B and C are disjoint then $P(B \cup C) = P(B) + P(C)$ (3) If B and C are not disjoint, we have the <i>inclusion-exclusion principle</i> $P(B \cup C) = P(B) + P(C) - P(B \cap C)$	The expected value or mean of X is defined by $E(X) = x_1 \cdot p(x_1) + x_2 \cdot p(x_2) + \dots + x_n \cdot p(x_n)$ $= \sum_{i=1}^n x_i \cdot p(x_i)$ $E(X + Y) = E(X) + E(Y)$ $E(aX + b) = aE(X) + b$ $E(h(X)) = \sum_{i=1}^n h(x_i) p(x_i)$
Multiplication theorem • For $A_1, A_2, \dots, A_n \in \mathcal{A}$ with $P(A_1 A_2 \dots A_n) > 0$ holds • For $P(B) > 0$ it holds $P(A \cap B) = P(A B) \cdot P(B)$ • For $P(A) > 0$ it holds $P(A \cap B) = P(B A) \cdot P(A)$	Probability mass function $p(a) = P(X = a)$ Cumulative distribution function $F(a) = P(X \leq a)$	Variance and standard deviation $\text{Var}(X) = E((X - \mu)^2) \quad \sigma = \sqrt{\text{Var}(X)}$ $= E(X^2) - \mu^2$ $\text{Var}(X) = \sum_{k=1}^n p(x_k) \cdot (x - \mu)^2$ $\text{Var}(X + b) = \text{Var}(X)$ $\text{Var}(aX) = a^2 \text{Var}(X)$ $\text{Var}(X) = 0$ then X is constant. X and Y be independent $\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$	
Independence $P(A B) = P(A) \quad P(A \cap B) = P(A) \cdot P(B)$ we say that the events A and B are independent .	Bernoulli distribution $Bernulli(p)$ X takes the values 0 and 1 $P(X = 1) = p$ and $P(X = 0) = 1 - p$ Binomial distribution $B(n, p)$ $p(k) = P(X = k) = \binom{n}{k} p^k \cdot (1 - p)^{n-k}$ Geometric distribution $p(k) = P(X = k) = (1 - p)^k \cdot p$ Poisson distribution $P(\lambda)$ $p(k) = P(X = k) = \frac{\lambda^k}{k!} \cdot e^{-\lambda}$ Hypergeometric $p(k) = P(X = k) = \frac{\binom{k}{r} \binom{N-k}{n-r}}{\binom{N}{n}}$ for $k = \max\{0, n - (N - r)\}, \dots, \min\{r, n\}$.	Expected value and variance $\mu = E(X) = \int_{-\infty}^{+\infty} x \cdot f(x) dx$ $\text{Var}(X) = E((X - E(X))^2)$ $= \int_{-\infty}^{+\infty} (x - \mu)^2 \cdot f(x) dx$ $\sigma = \sqrt{\text{Var}(X)}$	
Bayes' law Bayes' law based on the odds $\frac{P(H A)}{P(H^c A)} = \frac{P(H)}{P(H^c)} \cdot \frac{P(A H)}{P(A H^c)}$ a posteriori odds a priori odds likelihood quotient	Exponential distribution $\exp(\lambda)$ $f(x) = \lambda e^{-\lambda x}$, for $x \geq 0$. $F(x) = P(X \leq x) = \int_0^x \lambda e^{-\lambda t} dt = -e^{-\lambda t} \Big _0^x = 1 - e^{-\lambda x}$, $x \geq 0$ $E(X) = \frac{1}{\lambda} \quad \text{Var}(X) = \frac{1}{\lambda^2} \quad \sigma = \sqrt{\text{Var}(X)} = \frac{1}{\lambda}$ Uniform distribution $U(a, b)$ $f(x) = \begin{cases} \frac{1}{b-a}, & x \in (a, b) \\ 0, & \text{otherwise} \end{cases}$	Probability density function $P(c \leq x \leq d) = \int_c^d f(x) dx$ Cumulative distribution function $F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt$	
Independence $E(XY) = \int_a^b \int_c^d x \cdot y \cdot f(x, y) dx dy$ $E(XY) = \sum_{i=1}^n \sum_{j=1}^m x_i \cdot y_j \cdot p(x_i, y_j)$ Two random variables X and Y are independent if $E(XY) = E(X) \cdot E(Y)$	Gamma distribution $Gamma(\alpha, \beta)$ $f(x) = \frac{x^{\alpha-1} e^{-\frac{x}{\beta}}}{\Gamma(\alpha) \beta^\alpha}$, for $x > 0$. Gamma function $E(X) = \alpha\beta \quad \text{Var}(X) = \alpha\beta^2$ $\Gamma(\alpha) = \int_0^\infty u^{\alpha-1} e^{-u} du$ for $\alpha > 0$ $F(x) = P(X \leq x) = \int_0^x \frac{t^{\alpha-1} \cdot e^{-\frac{t}{\beta}}}{\Gamma(\alpha) \beta^\alpha} dt$, $x \geq 0$ $\chi^2(n)$ $P(X \leq \chi_{n,p}^2) = p$, for $p \in (0, 1)$ $f(x) = \frac{x^{\frac{n}{2}-1} \cdot e^{-\frac{x}{2}}}{\Gamma(\frac{n}{2}) \cdot 2^{\frac{n}{2}}}$ for $x > 0$.	Standardization $Z = \frac{X - \mu}{\sigma} \sim \mathcal{N}(0, 1)$ Kurtosis $k = \frac{E(X - E(X))^4}{\sigma^4}$ normal statistic $Z = \frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim \mathcal{N}(0, 1)$ • $k = 3$ - Normal • $k > 3$ - more "peaked" • $k < 3$ more "flat"	
Covariance $\text{Cov}(X, Y) = E((X - E(X)) \cdot (Y - E(Y)))$ • $\text{Cov}(aX + b, cY + d) = ac \text{Cov}(X, Y)$ for constants a, b, c and d . • $\text{Cov}(X_1 + X_2, Y) = \text{Cov}(X_1, Y) + \text{Cov}(X_2, Y)$ • $\text{Cov}(X, X) = \text{Var}(X)$ • $\text{Cov}(X, Y) = E(X \cdot Y) - E(X) \cdot E(Y)$ • $\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$ for any X and Y • If X and Y are independent, then $\text{Cov}(X, Y) = 0$.	Normal (Gaussian) distribution $\mathcal{N}(\mu, \sigma^2)$ $f(x) = \frac{1}{\sqrt{2\pi} \sigma^2} \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}}$, $\Phi(x) = \int_{-\infty}^x f(t) dt = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{t^2}{2}} dt$		
Correlation $-1 \leq \rho(X, Y) \leq 1$ $\rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)} \cdot \sqrt{\text{Var}(Y)}}$			

Confidence intervals for μ :

Normal z- statistic $Z = \frac{\bar{x}-\mu}{\frac{\sigma}{\sqrt{n}}} \sim \mathcal{N}(0,1)$ $Z = \frac{\bar{x}-\mu}{\frac{s}{\sqrt{n}}} \sim \mathcal{N}(0,1)$

Let X_1, \dots, X_n be a large sample from $X \sim \mathcal{N}(\mu, \sigma^2)$, with known σ

$$\left[\bar{x} - z_{\frac{\alpha}{2}} \cdot \frac{\sigma}{\sqrt{n}}, \bar{x} + z_{\frac{\alpha}{2}} \cdot \frac{\sigma}{\sqrt{n}} \right]$$

$$P(Z > z_{\frac{\alpha}{2}}) = \frac{\alpha}{2}, \quad z_{\frac{\alpha}{2}} = \Phi^{-1}\left(1 - \frac{\alpha}{2}\right)$$

Let X_1, \dots, X_n be a large sample from $\mathcal{N}(\mu, \sigma^2)$, with unknown σ

$$\left[\bar{x} - z_{\frac{\alpha}{2}} \cdot \frac{s}{\sqrt{n}}, \bar{x} + z_{\frac{\alpha}{2}} \cdot \frac{s}{\sqrt{n}} \right]$$

$$P(Z > z_{\frac{\alpha}{2}}) = \frac{\alpha}{2}, \quad z_{\frac{\alpha}{2}} = \Phi^{-1}\left(1 - \frac{\alpha}{2}\right)$$

Student's t- statistic

$$t = \frac{\bar{x}-\mu}{\frac{s}{\sqrt{n}}} \sim t_{n-1}$$

Let X_1, \dots, X_n be a small sample from $\mathcal{N}(\mu, \sigma^2)$, with unknown σ

$$\left[\bar{x} - t_{\frac{\alpha}{2}, n-1} \cdot \frac{s}{\sqrt{n}}, \bar{x} + t_{\frac{\alpha}{2}, n-1} \cdot \frac{s}{\sqrt{n}} \right]$$

$$P(t > t_{\frac{\alpha}{2}, n-1}) = \frac{\alpha}{2}$$

χ^2 -distribution

$$\chi^2 = \frac{s^2}{\sigma^2} = \frac{(n-1)s^2}{\sigma^2} \sim \chi_{n-1}^2$$

Let the data x_1, \dots, x_n is drawn from $\mathcal{N}(\mu, \sigma^2)$ with mean μ and standard deviation σ both unknown.

$$\frac{(n-1) \cdot s^2}{\chi_{\frac{\alpha}{2}, n-1}^2} \leq \sigma^2 \leq \frac{(n-1) \cdot s^2}{\chi_{1-\frac{\alpha}{2}, n-1}^2}$$

$$P(\chi^2 > \chi_{\alpha, n}^2) = \alpha$$

Small independent samples for $\mu_1 - \mu_2$: t-statistic

$$\left[(\bar{x}_1 - \bar{x}_2) - t^* \cdot s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}, (\bar{x}_1 - \bar{x}_2) + t^* \cdot s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} \right]$$

$$s_p = \sqrt{\frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1 + n_2 - 2}}$$

$$t^* = t_{n_1+n_2-2, \frac{\alpha}{2}}$$

$$t = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{s_p^2 \cdot \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$$

H_1	Rejection region	$P(t > t_{n_1+n_2-2, \alpha}) = \alpha$	$P(t > t_{n_1+n_2-2, \alpha}) = \alpha$	$P(t > t_{n_1+n_2-2, \frac{\alpha}{2}}) = \alpha$
$\mu_1 < \mu_2$	$t < -t_{n_1+n_2-2, \alpha}$			
$\mu_1 > \mu_2$	$t > t_{n_1+n_2-2, \alpha}$			
$\mu_1 \neq \mu_2$	$ t > t_{n_1+n_2-2, \frac{\alpha}{2}}$			

Sampling distribution of $\hat{p}_1 - \hat{p}_2$

$$\mu_{\hat{p}_1 - \hat{p}_2} = p_1 - p_2 \quad \text{and} \quad \sigma_{\hat{p}_1 - \hat{p}_2}^2 = \frac{p_1 q_1}{n_1} + \frac{p_2 q_2}{n_2}$$

$$\hat{p}_1 - \hat{p}_2 \sim \mathcal{N}\left(p_1 - p_2, \frac{p_1 q_1}{n_1} + \frac{p_2 q_2}{n_2}\right)$$

Confidence interval for $p_1 - p_2$

$$\left[(\hat{p}_1 - \hat{p}_2) - z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{\hat{p}_1 \hat{q}_1}{n_1} + \frac{\hat{p}_2 \hat{q}_2}{n_2}}, (\hat{p}_1 - \hat{p}_2) + z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{\hat{p}_1 \hat{q}_1}{n_1} + \frac{\hat{p}_2 \hat{q}_2}{n_2}} \right]$$

Hypothesis testing about $p_1 - p_2$

$$\sqrt{\frac{p_1 q_1}{n_1} + \frac{p_2 q_2}{n_2}} = \sqrt{p(1-p)\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}$$

$$z = \frac{(\hat{p}_1 - \hat{p}_2) - (p_1 - p_2)}{\frac{s_{p_1-p_2}}{\sqrt{n_1+n_2}}}$$

$$s_{p_1-p_2} = \sqrt{p(1-p)\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}$$

$$\bar{p} = \frac{x_1 + x_2}{n_1 + n_2} = \frac{n_1}{n_1 + n_2} \hat{p}_1 + \frac{n_2}{n_1 + n_2} \hat{p}_2$$

The F distribution

critical value $F_{\frac{\alpha}{2}, df1, df2} = F_{0.05, 12, 17}$

$$F = \frac{\text{larger sample variance}}{\text{smaller sample variance}} = \frac{s_1^2}{s_2^2}$$

One-way ANOVA Two estimates of σ^2

$$F = \frac{\text{variance between samples}}{\text{variance within samples}} = \frac{MSB}{MSW} \quad \text{the total sum of squares } SST = SSB + SSW =$$

the between-sample sum of squares SSB

$$SSB = \sum_i \frac{T_i^2}{n_i} - \frac{(\sum X)^2}{n}$$

the within-samples sum of squares SSW

$$SSW = \sum X^2 - \sum \frac{T_i^2}{n_i}$$

The variance between samples MSB and the variance within samples MSW are calculated as

$$MSB = \frac{SSB}{k-1} \quad MSW = \frac{SSW}{n-k}$$

where $k-1$ and $n-k$ are the df for the numerator and the df for the denominator for the F-distribution, and k is the number of samples.

x = the score of a student

k = the number of different samples

n_i = the size of sample i

T_i = the sum of the values in sample i

n = the number of values in all samples; $n = \sum_i n_i$

$\sum X$ = the sum of the values in all samples; $\sum X = \sum_i T_i$

$\sum X^2$ = the sum of the squares of the values in all samples

sampling distribution of \hat{p}

$$\hat{p} = \frac{X}{n} \quad X \sim B(n, p)$$

$$E(\hat{p}) = p \quad \text{and} \quad \text{Var}(\hat{p}) = \frac{pq}{n}$$

For large samples, the sampling distribution is approximately normal

$$\hat{p} \approx \mathcal{N}\left(p, \frac{pq}{n}\right)$$

$$\left[\hat{p} - z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{\hat{p}\hat{q}}{n}}, \hat{p} + z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{\hat{p}\hat{q}}{n}} \right] \quad \hat{p} = \frac{x}{n} \quad \text{and} \quad \hat{q} = 1 - \hat{p}$$

Adjusted confidence interval for p

$$\left[\hat{p} - z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{\hat{p}\hat{q}}{n+4}}, \hat{p} + z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{\hat{p}\hat{q}}{n+4}} \right] \quad \tilde{p} = \frac{x+2}{n+4} \quad \hat{q} = 1 - \tilde{p}$$

Polling confidence interval

Bernulli(p) distribution, where p is unknown

Conservative normal

$$\left[\bar{x} - z_{\frac{\alpha}{2}} \cdot \frac{1}{2\sqrt{n}}, \bar{x} + z_{\frac{\alpha}{2}} \cdot \frac{1}{2\sqrt{n}} \right]$$

Rule-of-thumb 95% confidence interval for p

$$\left[\bar{x} - \frac{1}{\sqrt{n}}, \bar{x} + \frac{1}{\sqrt{n}} \right]$$

Choosing a sample size

$$ME = z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}\hat{q}}{n}} \leq m \quad \text{margin of error}$$

Sampling distribution of $\bar{x}_1 - \bar{x}_2$

$$\mu_{\bar{x}_1 - \bar{x}_2} = \mu_1 - \mu_2 \quad \text{and} \quad \sigma_{\bar{x}_1 - \bar{x}_2}^2 = \sigma_{\bar{x}_1}^2 + \sigma_{\bar{x}_2}^2 = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$$

$$\bar{x}_1 - \bar{x}_2 \sim \mathcal{N}\left(\mu_1 - \mu_2, \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)$$

Confidence interval for $\mu_1 - \mu_2$

$$\left[(\bar{x}_1 - \bar{x}_2) - z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}, (\bar{x}_1 - \bar{x}_2) + z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} \right]$$

$$\left[(\bar{x}_1 - \bar{x}_2) - z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}, (\bar{x}_1 - \bar{x}_2) + z_{\frac{\alpha}{2}} \cdot \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}} \right]$$

Pooled sample estimator of σ^2

$$\sigma_1^2 = \sigma_2^2 = \sigma^2 \quad s_p^2 = \frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}$$

Pearson's Chi-square statistic goodness of fit

$$\chi^2 = \sum \frac{(O - E)^2}{E} \quad \text{degrees of freedom is } df = k - 1$$

O = observed frequency for a category

E = expected frequency for a category = np

O	p	$E = np$	$(O - E)^2$	$\frac{(O - E)^2}{E}$
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Test of independence $df = (R - 1) \cdot (C - 1)$

$$E = \frac{\text{(Row total)} \cdot \text{(Column total)}}{\text{Sample size}}$$

ANOVA Table

Source of Variation	Degrees of Freedom	Sum of Squares	Mean Square	Value of the Test Statistic
Between	$k - 1$	SSB	MSB	$F = \frac{MSB}{MSW}$
Within	$n - k$	SSW	MSW	
Total	$n - 1$	SST		

Regression and Correlation

$$r^2 = \frac{SSR}{SST} = \frac{\text{model's variation}}{\text{total variation}} = \text{"coefficient of determination"}$$

Analysis of variance (ANOVA)

$$F = \frac{MSR}{MSE} = \frac{SSR}{s^2}$$

Source of variation	SS	df	MS	F	p
Regression	SSR	1	$MSR = \frac{SSR}{1}$	$F = \frac{MSR}{MSE}$	p-value
Error	SST	$n - 2$	$MSE = \frac{SSE}{n - 2}$		
Total	SST	$n - 1$			

Test statistic

(t-statistic analog)

$$H_0: \mu = \mu_0$$

$$H_1: \mu < \mu_0$$

$$H_1: \mu > \mu_0$$

$$H_1: \mu \neq \mu_0$$

If the p -value $< \alpha$ we reject H_0

If the p -value $> \alpha$ we do not reject H_0

Decision

Truth	H_0 true H_1 true	fail to reject H_0	reject H_0
		\checkmark	Type 1 Error

$$z = \frac{\bar{x} - \mu_0}{\frac{\sigma}{\sqrt{n}}}$$

$$p\text{-value} = P(Z < z)$$

$$Z = \frac{\bar{x} - \mu}{\frac{s}{\sqrt{n}}}$$

$$p\text{-value} = P(Z > z)$$

$$p\text{-value} = P(|Z| > z)$$

	H_1	Rejection region	
$Z = \frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}}$	$\mu < \mu_0$	$(-\infty, -z_{\alpha})$	$P(Z > z_{\alpha}) = \alpha$
	$\mu > \mu_0$	$(z_{\alpha}, +\infty)$	$P(Z > z_{\alpha}) = \alpha$
	$\mu \neq \mu_0$	$(-\infty, z_{\frac{\alpha}{2}}) \cup (z_{\frac{\alpha}{2}}, +\infty)$	$P(Z > z_{\frac{\alpha}{2}}) = \alpha$
$t = \frac{\bar{x} - \mu}{\frac{s}{\sqrt{n}}}$	$\mu < \mu_0$	$(-\infty, -t_{n-1, \alpha})$	$P(t > t_{n-1, \alpha}) = \alpha$
	$\mu > \mu_0$	$(t_{n-1, \alpha}, +\infty)$	$P(t > t_{n-1, \alpha}) = \alpha$
	$\mu \neq \mu_0$	$(-\infty, t_{n-1, \frac{\alpha}{2}}) \cup (t_{\frac{\alpha}{2}}, +\infty)$	$P(t > t_{n-1, \frac{\alpha}{2}}) = \alpha$

Errors, significance level and power

Significance level = $P(\text{Type 1 Error})$

= probability we incorrectly reject H_0

= $P(\text{test statistic in rejection region} | H_0)$

$$= \alpha$$

Power = probability we correctly reject H_0

= $P(\text{test statistic in rejection region} | H_1)$

$$= 1 - P(\text{Type 2 Error})$$

$$= 1 - \beta$$

Hypothesis testing for a population proportion

$$z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$$

Test of hypothesis about a population variance

$$\chi^2 = \frac{(n-1)s^2}{\sigma_0^2} \sim \chi_{n-1}^2$$

H_1	Rejection region	
$\sigma^2 < \sigma_0^2$	$(0, \chi_{n-1, 1-\alpha}^2)$	$P(\chi^2 > \chi_{n-1, \alpha}^2) = 1 - \alpha$
$\sigma^2 > \sigma_0^2$	$(\chi_{n-1, \alpha}^2, +\infty)$	$P(\chi^2 > \chi_{n-1, \alpha}^2) = \alpha$
$\sigma^2 \neq \sigma_0^2$	$(0, \chi_{n-1, 1-\frac{\alpha}{2}}^2) \cup (\chi_{n-1, \frac{\alpha}{2}}^2, +\infty)$	$P(\chi^2 > \chi_{n-1, \frac{\alpha}{2}}^2) = \frac{\alpha}{2}$

Hypothesis testing about $\mu_1 - \mu_2$

$$z = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sigma_{\bar{x}_1 - \bar{x}_2}} \quad z = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}$$

H_1	Rejection region	
$\mu_1 < \mu_2$	$(-\infty, -z_{\alpha})$	$P(Z > z_{\alpha}) = \alpha$
$\mu_1 > \mu_2$	$(z_{\alpha}, +\infty)$	$P(Z > z_{\alpha}) = \alpha$
$\mu_1 \neq \mu_2$	$(-\infty, z_{\frac{\alpha}{2}}) \cup (z_{\frac{\alpha}{2}}, +\infty)$	$P(Z > z_{\frac{\alpha}{2}}) = \alpha$